

[Who Solved the Secretary Problem?]: Comment

Herbert Robbins

Statistical Science, Volume 4, Issue 3 (Aug., 1989), 291.

Your use of the JSTOR database indicates your acceptance of JSTOR's Terms and Conditions of Use. A copy of JSTOR's Terms and Conditions of Use is available at http://www.jstor.org/about/terms.html, by contacting JSTOR at jstor-info@umich.edu, or by calling JSTOR at (888)388-3574, (734)998-9101 or (FAX) (734)998-9113. No part of a JSTOR transmission may be copied, downloaded, stored, further transmitted, transferred, distributed, altered, or otherwise used, in any form or by any means, except: (1) one stored electronic and one paper copy of any article solely for your personal, non-commercial use, or (2) with prior written permission of JSTOR and the publisher of the article or other text.

Each copy of any part of a JSTOR transmission must contain the same copyright notice that appears on the screen or printed page of such transmission.

Statistical Science is published by Institute of Mathematical Statistics. Please contact the publisher for further permissions regarding the use of this work. Publisher contact information may be obtained at http://www.jstor.org/journals/ims.html.

Statistical Science
©1989 Institute of Mathematical Statistics

JSTOR and the JSTOR logo are trademarks of JSTOR, and are Registered in the U.S. Patent and Trademark Office. For more information on JSTOR contact jstor-info@umich.edu.

©2000 JSTOR

meet condition (b). Moreover, the upper bound for n = 3, j + 1 is (2 + k)/(4 + k) which is bigger than $\frac{1}{2}$, so condition (a) cannot be satisfied either.

Here is one more result for Pareto priors: If we reflect the one-sided priors (i.e., look at $-X_1, \dots, -X_n$), then the lower bound calculations are virtually the same as in the two-sided case, and the result is a slight improvement, to $(n-j)!/(\alpha+n)(\alpha+n-1)$ \dots $(\alpha+j+1)$ for any $\alpha>0$. For n=3, j=2, this is an improvement from 1/(k+4) to $1/(\alpha+3)$ —still not good enough to meet condition (b).

Thus the Ferguson Secretary Problem remains unsolved. Indeed, from these Pareto prior examples, it is not at all clear what the solution is: do the required exchangeable sequences exist or don't they? This quest for sufficiently "non-informative priors" should interest some Bayesians, too.

ADDITIONAL REFERENCE

HILL, B. M. (1968). Posterior distribution of percentiles: Bayes' theorem for sampling from a population. J. Amer. Statist. Assoc. 63 677-691.

Comment

Herbert Robbins

I am confused by Tom's attempt to clear up the confusion among various versions of the secretary problem. In Section 2 he defines the simplest form of the problem, in Section 4 he distinguishes secretary problems from Cayley's problem, etc. in which one observes numerical values of some possibly continuous random variable rather than just relative ranks, and in Section 5 he defines the 'general' secretary problem to be "a sequential observation and selection problem in which the payoff depends on the observations only through their relative ranks and not otherwise on their actual values." So far, so good. Then in Section 6 he introduces into the discussion the two-person googol game, which is not a secretary problem, and in Section 7 and Section 8 says that nobody has solved "the" secretary problem, possibly because no one realized that there was a game-theoretical problem to be solved. I can't agree with that.

Consider two cases of the secretary problem: (I) the payoff is 1 if we choose the best of the the n applicants, 0 otherwise, and we want to maximize the expected payoff, and (II) the loss is the absolute rank of the person selected (1 for the best, \cdots , n for the worst), and we want to minimize the expected loss. When all n! orders of the applicants are equally likely the solutions of (I) and (II) have been known and published for some time. And when the probabilities of the various permutations are controlled by an antagonist, so that (I) and (II) become game-theoretical (minimax) problems, their solutions are also in the litera-

ture: See problem 7 on page 60 of Chow, Robbins and Siegmund (1971), and page 89 of Chow, Moriguti, Robbins and Samuels (1964).

In the latter reference it is also shown that when the n! permutations are equally likely, the minimal expected loss for (II) with n applicants tends as $n \to \infty$ to the finite limit

$$A_1 = \prod_{i=1}^{\infty} \left(1 + \frac{2}{i}\right)^{1/(1+j)} \cong 3.8695.$$

This surprising result can be obtained by a heuristic argument involving a sequence of differential equations, but the argument is hard to make rigorous. The same heuristic argument yields a more general result: if the loss is taken to be $x(x+1) \cdots (x+k-1)$, where x is the absolute rank of the person selected and k is a fixed positive integer, then the minimal expected loss as $n \to \infty$ tends to

$$A_k = k! \left\{ \prod_{j=1}^{\infty} \left(1 + \frac{k+1}{j} \right)^{1/(k+j)} \right\}^k.$$

(As $k \to \infty$ the quantity in braces tends to $e^{\pi^2/6} \cong 5.1807$.) But when the loss is x^2 , rather than x or x(x+1), the limit as $n \to \infty$ of the minimal expected loss has not been exhibited explicitly by any formula such as this (it is, of course, less than A_2), nor has the minimax game-theoretical probability distribution of permutations been obtained for this case. Down with googol and up with problems like these!

ADDITIONAL REFERENCE

CHOW, Y. S., ROBBINS, H. and SIEGMUND, D. (1971). Great Expectations: The Theory of Optimal Stopping. Houghton Mifflin, Boston.

Herbert Robbins is Professor of Statistics at Rutgers University. His mailing address is: Hill Center for the Mathematical Sciences, Busch Campus, Rutgers University, New Brunswick, New Jersey 08903.